

Tilburg University

Identifiability in multiple time series

Tigelaar, H.H.

Publication date:
1977

[Link to publication in Tilburg University Research Portal](#)

Citation for published version (APA):

Tigelaar, H. H. (1977). *Identifiability in multiple time series*. (Ter discussie FEW; Vol. 77.066). Unknown Publisher.

General rights

Copyright and moral rights for the publications made accessible in the public portal are retained by the authors and/or other copyright owners and it is a condition of accessing publications that users recognise and abide by the legal requirements associated with these rights.

- Users may download and print one copy of any publication from the public portal for the purpose of private study or research.
- You may not further distribute the material or use it for any profit-making activity or commercial gain
- You may freely distribute the URL identifying the publication in the public portal

Take down policy

If you believe that this document breaches copyright please contact us providing details, and we will remove access to the work immediately and investigate your claim.

R
CBM
R
1977-66
7627
1977
66



KATHOLIEKE HOGESCHOOL TILBURG

| | | |
|---|---|--|
| Bestemming  | TIJDSCHRIFTENBUREAU BIBLIOTHEEK KATHOLIEKE HOGESCHOOL TILBURG | Nr.  |
|---|---|--|

REEKS "TER DISCUSSIE"





KATHOLIEKE HOGESCHOOL TILBURG

"REEKS TER DISCUSSIE"

No 77.066

Oktober 1977

Identifiability in Multiple Time Series

Drs. H.H. Tigelaar.

R11

T Time Series

SUBFACULTEIT ECONOMETRIE

SUMMARY.

This paper deals with identifiability of matrixcoefficients in multivariate stochastic difference-equation-models, where identifiability should be interpreted as in [5], that is, we are only concerned with identifiability with respect to the class $\mathcal{P}^{(n)}$, the class of possible distributions of a sample taken from the observable process at n consecutive time points. (For a discussion on this: see [2] and [5]).

CONTENTS.

| | | |
|--------------------------------------|-----------|----|
| 1. Introduction | - - - - - | 1 |
| 2. Notations and preliminary results | - - - - - | 2 |
| 3. The basic lemma | - - - - - | 5 |
| 4. The AR-case | - - - - - | 10 |
| 5. The ARMA-case | - - - - - | 12 |
| 6. References | - - - - - | 18 |

1. Introduction.

In this paper we study m -variate stochastic processes with complex-valued components and finite second moments, which are assumed to be the unique (a.s.) weakly stationary solution of a difference-equation of the form:

$$(1.1) \quad \sum_{k=0}^p A_k x_{t-k} = \sum_{j=0}^q B_j \varepsilon_{t-j} \quad t = 0, \pm 1, \pm 2, \dots$$

where $\{\varepsilon_t\}$ is (unobservable) m -variate white noise, with unknown covariance-matrix Σ_ε . Furthermore, the $m \times m$ matrices A_k , $k = 1, \dots, p$ and B_j , $j = 1, \dots, q$ are supposed to be unknown. As in [5] we introduce a parameter ξ , characterizing the distribution of the process $\{\varepsilon_t\}$, and we put:

$$\theta = (A_1, \dots, A_p \quad ; \quad B_1, \dots, B_q, \xi)$$

In section 3 identifiability will be proved for

$$\psi(\theta) = (A_1, \dots, A_p \quad , \quad B_1, \dots, B_q, \Sigma_\varepsilon(\xi))$$

for the special case $p = 0$ (Moving Average).

In section 5 the general case will be treated.

2. Notations and preliminary results.

Let $\{x_t\}$ be a m -variate weakly stationary process. Then its matrix-valued spectral measure (see: [3]) will be denoted by F_x and its spectral representation by:

$$\int e^{it\lambda} \underline{z}_x \{d\lambda\}$$

where $\underline{z}_x(\lambda)$ is the associated m -variate process with orthogonal increments with:

$$E\{\underline{x}_x\{d\lambda\} \underline{z}_x^*\{d\lambda\}\} = F_x\{d\lambda\}.$$

(The $*$ indicates the complex conjugate transpose). Furthermore we introduce the real-valued measure:

$$\Phi_x\{d\lambda\} = \text{tr } F_x\{d\lambda\}$$

It is easily seen that $\Phi_x\{D\} = 0$ i.f.f. (if and only if) $F_x\{D\} = 0$. Thus we can write:

$$F_x\{d\lambda\} = \tilde{f}_x(\lambda) \Phi_x\{d\lambda\}.$$

When the measure Φ_x is absolutely continuous with respect to Lebesgue measure, (with density $\varphi_x(\lambda)$), F_x is said to be so and we can write:

$$F_x\{d\lambda\} = f_x(\lambda) d\lambda$$

where $f_x(\lambda) = \tilde{f}_x(\lambda) \varphi_x(\lambda)$ is called the spectral density matrix.

When we are dealing with processes, satisfying the difference equation (1.1), it is usefull to define the matrix-generating functions:

$$A(z) = \sum_{k=0}^p A_k z^k,$$

$$B(z) = \sum_{k=0}^q B_k z^k.$$

To achieve maximum analogy between the univariate and the multivariate case, we introduce the concept of a singularity-point:

Definition 2.1.

Let $Q(z)$ be a square matrix-valued function of the complex variable z . Then the point z_0 is called a singularity point of $Q(z)$ if $\det Q(z_0) = 0$.

Definition 2.2.

A matrixvalued function of the form $\sum_{k=0}^p A_k z^k$ is called a matrixpolynomial;

if $A_p \neq 0$ p is the degree of the matrixpolynomial.

It will be clear now, that singularitypoints of matrixpolynomials will take over the role of the zeros of polynomials in the univariate case.

Definition 2.3.

The matrixfunction $Q(z)$ is called analytic (or holomorphic), if all its elements are analytic functions of z .

Since singularity points of an analytic matrix $Q(z)$ are in fact zeros of the analytic function $\det Q(z)$, they are either isolated points, and hence the set of singularity points has lebesgue-measure 0, or $\det Q(z)$ is identically zero.

We can now state:

Theorem 2.1.

The homogeneous difference-equation

$$(2.1.) \quad \sum_{k=0}^p A_k x_{t-k} = 0 \quad t = 0, \pm 1, \pm 2, \dots$$

has a non-trivial weakly stationary solution i.f.f. $A(e^{-i\lambda})$ has as a function of λ at least one singularitypoint. Furthermore, if S is the set of singularity points and S' its complement, we have for all solutions $\{x_t\}$:

$$F_x \{S'\} = 0$$

Proof. Suppose:

$$\underline{x}_t = \int e^{it\lambda} \underline{z}_x\{d\lambda\}$$

is a weakly stationary solution of (2.1), thus:

$$\int A(e^{-i\lambda}) e^{it\lambda} \underline{z}_x\{d\lambda\} \stackrel{a.s.}{=} 0 \Leftrightarrow$$

$$\int A(e^{-i\lambda}) F_x\{d\lambda\} A^*(e^{-i\lambda}) = 0 \Leftrightarrow$$

$$\int A(e^{-i\lambda}) \tilde{f}_x(\lambda) A^*(e^{-i\lambda}) \phi_x\{d\lambda\} = 0 \Leftrightarrow$$

$$A(e^{-i\lambda}) \tilde{f}_x(\lambda) A^*(e^{-i\lambda}) \stackrel{a.e.}{=} 0, \text{ w.r.t. } \phi_x \text{ (or } F_x).$$

Hence, for $\lambda \in S'$ we must have:

$$\tilde{f}_x(\lambda) \stackrel{a.e.}{=} 0 \text{ w.r.t. } \phi_x \Rightarrow$$

$$F_x\{S'\} = \int_{S'} F_x\{d\lambda\} = \int_{S'} \tilde{f}_x(\lambda) \phi_x\{d\lambda\} = 0$$

Remark: In contrast to the univariate case, the m-variate homogeneous difference equation may have a non trivial solution with absolutely continuous spectrum, due to the fact that $\det A(e^{-i\lambda})$ can be identically zero.

3. The basic lemma.

Lemma 3.1.

Let $Q(z)$ be a $m \times m$ matrixpolynomial of degree $p > 0$, which is non-singular for at least one point z_1 , and let Σ be an arbitrary hermitian positive definite $m \times m$ matrix.

Then there exists a matrixpolynomial $Q_0(z)$ with:

- a) $Q_0(z)$ nonsingular for $|z| < 1$.
- b) $Q(e^{-i\lambda}) \Sigma Q^*(e^{-i\lambda}) = Q_0(e^{-i\lambda}) \Sigma Q_0^*(e^{-i\lambda}) \quad \forall \lambda \in [-\pi, \pi]$

Furthermore $Q_0(z)$ is of degree p and is uniquely determined except for a constant $m \times m$ matrix H , with $H \Sigma H^* = \Sigma$.

Proof. Since Σ can be decomposed into TT^* , and the matrix T can be absorbed into $Q(z)$, it is no restriction to take $\Sigma = I_m$, the $m \times m$ unit-matrix.

► Existence.

Let z_0 be a singularitypoint of $Q(z)$ with $|z_0| < 1$, and let c_1, \dots, c_k be an orthonormal basis for $\ker Q(z_0)$. The matrix with columns c_1, \dots, c_k is denoted by C .

Let d_1, \dots, d_{m-k} be an orthonormal basis for $[\ker Q(z_0)]^\perp$, the orthogonal complement of $\ker Q(z_0)$ and denote the matrix with columns d_1, \dots, d_{m-k} by D . Put:

$$U = [C; D]$$

then we have:

$$U^*U = UU^* = I_m.$$

Since:

$$Q(z_0)C = 0$$

there exists an integer $a_0 > 0$, such that:

$$Q(z) C = (z-z_0)^{a_0} Q_1(z) \quad \forall z,$$

where $Q_1(z)$ is a $m \times m$ matrix-polynomial of degree $p-a_0$ and $\text{rank}(Q_1(z_0)) = k$.

Thus:

$$Q(z)U = [(z-z_0)^{a_0} Q_1(z); Q(z)D].$$

Post-multiplying by U^* and using (2.2) yields:

$$Q(z) = [(z-z_0)^{a_0} Q_1(z); Q(z)D] U^*$$

Put:

$$\tilde{Q}(z) = [(z\bar{z}_0-1)^{a_0} Q_1(z); Q(z)D] U^*$$

Then we have:

$$\begin{aligned} \det \tilde{Q}(z) &= (z\bar{z}_0-1)^{a_0 k} \det [Q_1(z); Q(z)D] \det U^* \\ &= \left(\frac{z\bar{z}_0-1}{z-z_0} \right)^{a_0 k} \det Q(z). \end{aligned}$$

Thus, ignoring multilicities, $\tilde{Q}(z)$ has at least one singularity point less than $Q(z)$ in $|z| < 1$. Furthermore:

$$\begin{aligned} \tilde{Q}(e^{-i\lambda}) \tilde{Q}^*(e^{-i\lambda}) &= |e^{-i\lambda} \bar{z}_0 - 1|^{2a_0} Q_1(e^{-i\lambda}) Q_1^*(e^{-i\lambda}) + Q(e^{-i\lambda}) D D^* Q^*(e^{-i\lambda}) \\ &= |e^{-i\lambda} - z_0|^{2a_0} Q_1(e^{-i\lambda}) Q_1^*(e^{-i\lambda}) + Q(e^{-i\lambda}) D D^* Q^*(e^{-i\lambda}) \\ &= Q(e^{-i\lambda}) Q^*(e^{-i\lambda}), \end{aligned}$$

so that $\tilde{Q}(z)$ satisfies condition b). Since this procedure can be repeated for every singularity point, we can find in finitely many steps a matrix $Q_0(z)$ of degree p satisfying a) and b).

► Uniqueness.

Suppose we have two matrixpolynomials Q_0 and P_0 , both of degree p , satisfying a) and b). If z_0 is a singularitypoint of Q_0 and P_0 with $|z_0| = 1$, then b) implies:

$$\ker Q_0^*(z_0) = \ker P_0^*(z_0).$$

Thus we can write as before:

$$Q_0^*(z) = [((z-z_0)^{a_0} Q_1(z))^*, Q_0^*(z)D] U^*$$

$$P_0^*(z) = [((z-z_0)^{b_0} P_1(z))^*, P_0^*(z)D] U^*$$

where $\det [Q_1^*(z_0), Q_0^*(z_0)D] \neq 0$ and

$$\det [P_1^*(z_0), P_0^*(z_0)D] \neq 0$$

Since b) implies:

$$|\det Q_0^*(e^{-i\lambda})| = |\det P_0^*(e^{-i\lambda})|, \quad \forall \lambda$$

it is easily seen that we must have $a_0 = b_0$.

Therefore it suffices to prove the uniqueness of:

$$W(z) = \begin{bmatrix} Q_1(z) \\ D^* Q_0(z) \end{bmatrix}$$

Put:

$$T(z) = \begin{bmatrix} P_1(z) \\ D^* P_0(z) \end{bmatrix}$$

Then we have:

$$Q(z) = \mathcal{V}(z) W(z), \quad P_0(z) = \mathcal{V}(z) T(z)$$

where:

$$V(z) = \begin{bmatrix} (z-z_0)^{a_0} I_k & & 0 \\ & \ddots & \\ 0 & & I_{m-k} \end{bmatrix}$$

Thus for all z we have:

$$W(z) = \lim_{\eta \rightarrow z} V(\eta)^{-1} Q_0(z), \quad T(z) = \lim_{\xi \rightarrow z} V(\xi)^{-1} P_0(z).$$

Hence:

$$\begin{aligned} W(e^{-i\lambda}) W^*(e^{-i\lambda}) &= \lim_{\eta \rightarrow e^{-i\lambda}} V(\eta)^{-1} Q_0(e^{-i\lambda}) Q_0^*(e^{-i\lambda}) V(\eta)^{-1*} = \\ &= \lim_{\xi \rightarrow e^{-i\lambda}} V(\xi)^{-1} P_0(e^{-i\lambda}) P_0^*(e^{-i\lambda}) V(\xi)^{-1*} \\ &= T(e^{-i\lambda}) T^*(e^{-i\lambda}). \end{aligned}$$

Thus $W(z)$ and $T(z)$ also satisfy b), but have at least one singularity point less than $P_0(z)$ and $Q_0(z)$ on $|z| = 1$.

In this way we can remove all singularity points from the unit circle, and therefore it is no restriction to assume $Q_0(z)$ and $P_0(z)$ non-singular on $|z| = 1$.

Consider:

$$V(z) = P_0^{-1}(z) Q_0(z)$$

Using b) we have:

$$\begin{aligned} V(e^{-i\lambda}) V^*(e^{-i\lambda}) &= P_0^{-1}(e^{-i\lambda}) Q_0(e^{-i\lambda}) Q_0^*(e^{-i\lambda}) P_0^{*-1}(e^{-i\lambda}) = \\ &= P_0^{-1}(e^{-i\lambda}) P_0(e^{-i\lambda}) P_0^*(e^{-i\lambda}) P_0^{*-1}(e^{-i\lambda}) = I_m \end{aligned}$$

or equivalently: $V(z)$ is unitary on $|z| = 1$. Furthermore $V(z)$ is analytic inside the unit circle and has no singularity points on $|z| \leq 1$.

Therefore both $V(z)$ and $V^{-1}(z)$ can be expanded into a power series:

$$V(z) = \sum_{k=0}^{\infty} M_k z^k, \quad V^{-1}(z) = \sum_{k=0}^{\infty} N_k z^k, \quad |z| < 1.$$

Since the elements of $V(z)$ and $V^{-1}(z)$ are rational functions with no poles on $|z| \leq 1$, both series converge on the unitcircle.

Hence:

$$V^{-1}(e^{-i\lambda}) = \sum_{k=0}^{\infty} N_k e^{-ik\lambda} = V^*(e^{-i\lambda}) = \sum_{k=0}^{\infty} M_k^* e^{ik\lambda} \Rightarrow$$

$$\Rightarrow \begin{cases} N_0 = M_0^* \\ N_k = M_k = 0, k \geq 1 \end{cases}$$

by a simple equation of coefficients.

It follows that $V(z) = N_0$, or equivalently:

$$Q_0(z) = P_0(z) N_0, \text{ with } N_0 \text{ unitary.}$$

This proves the lemma.

As an immediate consequence of the lemma we have:

Theorem 3.1.

When in the m -variate Moving Average case the parameterspace θ is such that $B(z)$ has no singularitypoints in $|z| < 1$, and $B_0 = I_m$, then $\Psi(\theta) = (B_1, \dots, B_q, \Sigma_\epsilon)$ is identifiable w.r.t. $\mathcal{P}^{(q+1)}$.

The proof is exactly that of the univariate case (see [5]) and will be omitted.

Remark: The distributions of the class $\mathcal{P}^{(q+1)}$ are in fact $m(q+1)$ -dimensional distributions, and there are $m^2 q + \frac{1}{2} m(m+1)$ unknown parameters. Thus for $m > 1$ one has identification with less observations than unknown parameters.

4. The AR-case.

Theorem 4.1.

When in the AR-case θ is such that $A_0 = I_m$, $A(z)$ has no singularity points on $|z| \leq 1$, and Σ_ϵ is non-singular, then:

$$\Psi(\theta) = (A_1 \dots A_p, \Sigma_\epsilon)$$

is identifiable w.r.t. $\mathcal{P}^{(p+1)}$.

Proof.

As in the univariate case, we may write down the YULE-WALKER-equations:^{*})

$$(4.1) \quad \Psi(\theta) \mathcal{R} = -R$$

where

$$R = (\Gamma_0, \dots, \Gamma_p), \quad \Gamma_s = E\{\underline{x}_t \underline{x}_{t-s}^*\}$$

and

$$\mathcal{R} = \begin{bmatrix} \Gamma_1^* & \Gamma_0 & \dots & \Gamma_{p-1} \\ \Gamma_2^* & \Gamma_1^* & \dots & \Gamma_{p-2} \\ \dots & \dots & \dots & \dots \\ \Gamma_r^* & \Gamma_{r-1}^* & \dots & \Gamma_0 \\ -I_m & 0 & \dots & 0 \end{bmatrix}$$

Thus we have identifiability when there is no non-trivial linear combination

$$\sum_{j=1}^p c_j^* \underline{x}_{t-j}, \quad c_j \in \mathbb{C}^m$$

with variance zero.

^{*}) The condition that $A(z)$ has no singularity-points on $|z| \leq 1$ implies the existence of a non-sided MA-representation which in turn implies (4.1)

Suppose

$$(4.2) \quad \sum_{j=1}^p c_j^! \underline{x}_{t-j} \stackrel{\text{a.s.}}{=} 0$$

Since the process $\{\underline{x}_t\}$ has a spectral density given by:

$$f_x(\lambda) = A^{-1}(e^{-i\lambda}) \sum_{\epsilon} A^{-1*}(e^{-i\lambda})$$

we must have:

$$\sum_{j=1}^p \int c_j^! A^{-1}(e^{-i\lambda}) \sum_{\epsilon} A^{-1*}(e^{-i\lambda}) c_j^! d\lambda = 0$$

But then it follows easily that $c_j = 0$ $j = 1, \dots, p$ since all terms are real and nonnegative.

Note that (4.2) does not imply a singular spectral measure (see § 2).

5. The ARMA-case.

The multivariate ARMA-case needs special attention. The generalization from the univariate case is far from obvious: matrixpolynomials don't factorize as simple as polynomials do, and there is no 1-1-correspondence between singularitypoints and factorizations; besides singularitypoints, also the corresponding nullspaces play an important role.

The following terminology will be useful:

Definition 5.1.

The square matrices A and B are called comparable when they have the same nullspaces. They are called completely incomparable when the nullspaces have null-intersection.

HANNAN, who proved the identifiability with respect to the class $\mathcal{P}^{(\infty)}$ in [4],

showed that a necessary condition for the matrixpolynomials $A(z) = \sum_{k=0}^p A_k z^k$ and $B(z) = \sum_{k=0}^q B_k z^k$ to have no common left factor^{*}, is that A_p^* and B_p^* are

completely incomparable, and he achieves identifiability by the condition that $A(z)$ and $B(z)$ have the unitmatrix as a greatest common left-divisor (g.c.l.d) (see mc. Duffee, [1] p 35).

The last condition is difficult to verify and a condition in terms of singularitypoints and nullspaces looks more appropriate.

The following lemma shows what kind of condition will be needed.

Lemma 5.1.

When the $m \times m$ -matrix polynomials $A(z)$ and $B(z)$ have a common singularitypoint z_0 , and $A^*(z_0)$ and $B^*(z_0)$ are not completely incomparable, then $A(z)$ and $B(z)$ have a common left factor.

* Obviously this is necessary for identifiability when only second-order properties are considered.

Proof:

Let C be the $m \times c$ matrix with columns an orthogonal basis for the intersection of the nullspaces of $A^*(z_0)$ and $B^*(z_0)$, and let D be the matrix with columns an orthonormal basis for the orthogonal complement.

In exactly the same way as in the proof of the basic lemma 3.1. we have:

$$U^* A(z) = \begin{bmatrix} (z-z_0)A_1(z) \\ D^* A(z) \end{bmatrix}$$

$$U^* B(z) = \begin{bmatrix} (z-z_0)B_1(z) \\ D^* B(z) \end{bmatrix}$$

where $U = [C \ ; \ D]$, $UU^* = U^*U = I_m$.

Hence:

$$A(z) = U \begin{bmatrix} (z-z_0)I_c & 0 \\ 0 & I_{m-c} \end{bmatrix} \begin{bmatrix} A_1(z) \\ D^* A(z) \end{bmatrix}$$

$$B(z) = U \begin{bmatrix} (z-z_0)I_c & 0 \\ 0 & I_{m-c} \end{bmatrix} \begin{bmatrix} B_1(z) \\ D^* B(z) \end{bmatrix}$$

which proves the lemma.

We now have:

Theorem 5.2.

When in the ARMA-case θ is such that:

- a) $A_0 = B_0 = I_m$,
- b) $A(z)$ has no singularitypoints on $|z| \leq 1$
 $B(z)$ has no singularitypoints on $|z| < 1$
- c) $A^*(z)$ and $B^*(z)$ are completely incomparable \forall_z
 A_p^* and B_q^* are completely incomparable.

then:

$$\Psi(\theta) = (A_1, \dots, A_p, B_1, \dots, B_q, \Sigma_\epsilon)$$

is identifiable w.r.t. $\mathcal{P}^{(p+q+1)}$.

Proof.

Put:

$$\Psi_1(\theta) = (A_1, \dots, A_p), \Psi_2(\theta) = (B_1, \dots, B_q, \Sigma_\epsilon).$$

Postmultiplying both sides of the difference-equation (1.1) with \underline{x}_{t-s}^* , $s = q+1, q+2, \dots$ and taking expectations yields:

$$(5.1) \quad \sum_{k=0}^p A_k \Gamma_{s-k} = 0, \quad s = q+1, q+2, \dots \quad *)$$

where

$$\Gamma_s = E\{\underline{x}_t \underline{x}_{t-s}^*\}.$$

Since $A_0 = I_m$ this can for $s = q+1, \dots, q+p$ be written:

$$\Psi_1(\theta) \mathcal{R}_q = -R',$$

where

$$R = \begin{bmatrix} \Gamma_{q+1} \\ \vdots \\ \Gamma_{q+p} \end{bmatrix} \quad \text{and} \quad \mathcal{R}_q = \begin{bmatrix} \Gamma_q & \Gamma_{q+1} & \dots & \Gamma_{q+p-1} \\ \vdots & \vdots & \ddots & \vdots \\ \Gamma_{q-1} & \Gamma_q & \dots & \Gamma_{q+p-2} \\ \vdots & \vdots & \ddots & \vdots \\ \Gamma_{q-p+1} & \Gamma_{q-p+2} & \dots & \Gamma_q \end{bmatrix}$$

Suppose \mathcal{R}_q is singular; then there exist vectors c_0, \dots, c_{p-1} , not all zero, such that:

*) The condition that $A(z)$ has no singularity points on $|z| \leq 1$ implies that there exists a one-sided moving average representation, which in turn implies (5.1).

$$(5.2) \quad \sum_{s=0}^{p-1} \Gamma_{j+s} c_s = 0 \quad j = q-p+1, q-p+2, \dots, q$$

From (5.1) we have:

$$\Gamma_s = - \sum_{k=1}^p A_k \Gamma_{s-k} \quad s = q+1, q+2, \dots$$

or equivalently:

$$\Gamma_{q+1+s} = - \sum_{k=1}^p A_k \Gamma_{q+1+s-k} \quad s = 0, 1, \dots, p-1$$

Postmultiplying by c_s and summing over s yields:

$$\sum_{s=0}^{p-1} \Gamma_{q+1+s} c_s = - \sum_{k=1}^p A_k \sum_{s=0}^{p-1} \Gamma_{q+1+s-k} c_s = 0$$

using (5.2). Continuing this way (5.2) is easily seen to be valid for $j \geq q-p+1$.

Let

$$y_t = \sum_{k=0}^{p-1} c_k^* x_{t-k} \quad t = 0, \pm 1, \pm 2, \dots$$

and

$$c(z) = \sum_{k=0}^{p-1} c_k^* z^k.$$

Then $\{y_t\}$ is a weakly stationary process with spectral density matrix:

$$\frac{1}{2\pi} c(e^{-i\lambda}) A^{-1}(e^{-i\lambda}) B(e^{-i\lambda}) \Sigma_{\epsilon} B^*(e^{-i\lambda}) A^{-1*}(e^{-i\lambda}) c^*(e^{-i\lambda})$$

(see: [3] § II.4).

(Here we use the fact that Σ_{ϵ} is nonsingular; otherwise y_t could be identically zero.)

$$E\{x_{t-j} y_{t-s}^*\} = \sum_{k=1}^{p-1} \Gamma_{s-j+k} c_k = 0 \quad , \quad s_j \geq q-p+1$$

Premultiplying with c_j^* , and summing over j yields:

$$E\{y_t y_{t+s}^*\} = 0 \quad , \quad |s| \geq q$$

This can equivalently be written as:

$$\int e^{is\lambda} c(e^{-i\lambda}) A^{-1}(e^{-i\lambda}) B(e^{-i\lambda}) \Sigma_{\epsilon} B^*(e^{-i\lambda}) A^{-1*}(e^{-i\lambda}) c^*(e^{-i\lambda}) d\lambda = 0, \quad |s| \geq q$$

But then it follows that $c(z)A^{-1}(z)B(z)$ must be a matrixpolynomial whose degree is at most $q-1$.

Let

$$K(z) = (\det A(z)) A^{-1}(z)$$

then also $K(z)$ is a matrixpolynomial in z with:

$$\det K(z) = (\det A(z))^{m-1}$$

Denote the degree of $c(z)K(z)$ by s and the degree of $\det A(z)$ by a .

Since $c(z)A^{-1}(z)B(z)$ is a polynomial of degree at most $q-1$, say $q-r$, and $\det A(z)$ is a scalar polynomial, we must have:

$$s+q-a \geq q-r$$

thus:

$$a \leq s+r, \quad r \geq 1$$

1) Suppose: $a \leq s$.

Put:

$$s(z) = c(z) K(z) = \sum_{j=0}^s s_j z^j$$

Since $s(z)B(z)$ is a matrixpolynomial of degree at most $q-1+a$ we must have:

$$\sum_{k=0}^s s_k B_{1-k} = 0, \quad 1 \geq q+a$$

In the same way for $s(z) A(z) = c(z) \det A(z)$ we must have:

$$\sum_{k=0}^s s_k A_{n-k} = 0, \quad n \geq p+a$$

Choosing $l = q+s$ and $n = p+s$ yields:

$$s_{s_q} B_q = s_{s_p} A_p = 0$$

which contradicts the complete incomparability of A_p^* and B_q^* .

2.) When $a \geq s+1$, it is easily seen that there exists at least one zero, say z_0 of $\det A(z)$ with:

$$s(z_0) \neq 0$$

$$s(z) B(z) = (z-z_0) V(z),$$

for some polynomial $V(z)$.

Thus $B(z_0)$ is singular and $B^*(z_0) s^*(z_0) = 0$.

On the other hand we have:

$$s(z_0) A(z_0) = c(z_0) \det A(z_0) = 0,$$

thus also $A^*(z_0) s^*(z_0) = 0$, which contradicts the complete incomparability of $A(z)$ and $B(z)$ for all z , and completes the proof of the identifiability of $\Psi_1(\theta)$.

We now fix $\Psi_1(\theta)$ and put:

$$\underline{z}_t = \sum_{k=0}^p A_k \underline{x}_{t-k} = \sum_{j=0}^q B_j \underline{\varepsilon}_{t-j}$$

Thus $\{\underline{z}_t\}$ is a MA-process, with distribution independent of $\Psi_1(\theta)$. If $\tilde{P}_{\Psi_2(\theta)}^{(n)}$ denotes the joint distribution of $\underline{z}_1, \dots, \underline{z}_n$, and

$$\tilde{\mathcal{P}}^{(n)} = \{\tilde{P}_{\Psi_2(\theta)}^{(n)} \mid \theta \in \Theta\},$$

Then we have the implication:

$$\tilde{P}_{\Psi_2(\theta_1)}^{(n)} \neq \tilde{P}_{\Psi_2(\theta_2)}^{(n)} \Rightarrow P_{\theta_1}^{(n+p)} \neq P_{\theta_2}^{(n+p)}$$

Thus since $\Psi_2(\theta)$ is identifiable w.r.t. $\tilde{\mathcal{P}}^{(q+1)}$ (according to theorem 3.1), we also have identifiability w.r.t. $\tilde{\mathcal{P}}^{(p+q+1)}$. This proves the theorem.

REFERENCES.

- 1 MAC DUFFEE, C.C., (1956). The Theory of matrices. New York Chelsea.
- 2 GENUGTEN, B.v.d., Identifiability in Statistical inference.
Statistica Neerlandica.
- 3 HANNAN, E.J., (1970) Multiple Time Series. Wiley, New York.
- 4 HANNAN, E.J., (1969) The identification of vector mixed autoregressive-
moving average systems.
Biometrika 56, 223-225.
- 5 TIGELAAR, H.H., (1976) Identifiability in models with lagged variables
Reeks "Ter Diskussie" 76.020
Catholic University, Tilburg.

In de Reeks ter Discussie zijn verschenen:

| | | |
|--|--|---------------|
| 1.H.H. Tigelaar | Spectraalanalyse en stochastische lineaire differentievergelijkingen. | juni '75 |
| 2.J.P.C.Kleijnen | De rol van simulatie in de algemene econometrie. | juni '75 |
| 3.J.J. Kriens | A stratification procedure for typical auditing problems. | juni '75 |
| 4.L.R.J. Westermann | On bounds for Eigenvalues | juni '75 |
| 5.W. van Hulst J.Th. van Lieshout | Investment/financial planning with endogenous lifetimes: a heuristic approach to mixed integer programming. | juli '75 |
| 6.M.H.C.Paardekooper | Distribution of errors among input and output variables. | augustus '75 |
| 7.J.P.C. Kleijnen | Design and analysis of simulation Practical statistical techniques. | augustus '75 |
| 8.J. Kriens | Accountantscontrole met behulp van steekproeven. | september '75 |
| 9.L.R.J. Westermann | A note on the regula falsi | september '75 |
| 10.B.C.J. van Velthoven | Analoge simulatie van economische modellen. | november '75 |
| 11.J.P.C. Kleijnen | Het economisch nut van nauwkeurige informatie: simulatie van ondernemingsbeslissingen en informatie. | november '75 |
| 12.F.J. Vandamme | Theory change, incompatibility and non-deductibility. | december '75 |
| 13.A. van Schaik | De arbeidswaardeleer onderbouwd? | januari '76 |
| 14.J.vanLieshout J.Ritzen J.Roemen | Input-ouputanalyse en gelaagde planning. | februari '76 |
| 15.J.P.C.Kleijnen | Robustness of multiple ranking procedures: a Monte Carlo experiment illustrating design and analysis techniques. | februari '76 |
| 16.J.P.C. Kleijnen | Computers and operations research: a survey. | februari '76 |
| 17.J.P.C. Kleijnen | Statistical problems in the simulation of computer systems. | april '76 |
| 18.F.J. Vandamme | Towards a more natural deontic logic. | mei '76 |
| 19.J.P.C. Kleijnen | Design and analysis of simulation: practical, statistical techniques. | juni '76 |
| 20.H.H. Tigelaar | Identifiability in models with lagged variables. | juli '76 |
| 21.J.P.C. Kleijnen | Quantile estimation in regenerative simulation: a case study. | augustus '76 |
| 22.W.Derks | Inleiding tot econometrische modellen van landen van de E.E.G. | augustus '76 |
| 23.B. Diederer Th. Reijs W. Derks | Econometrisch model van België. | september '76 |
| 24.J.P.C. Kleijnen | Principles of Economics for computers. | augustus '76 |
| 25.B. van Velthoven | Hybriede simulatie van economische modellen. | augustus '76. |

| | | |
|-------------------------------------|--|---------------|
| 26. F. Cole | Forecasting by exponential smoothing, the Box and Jenkins procedure and spectral analysis. A simulation study. | september '76 |
| 27. R. Heuts | Some reformulations and extensions in the univariate Box-Jenkins time series analysis. | juli '76 |
| 28. W. Derks | Vier econometrische modellen. | |
| 29. J. Frijns | Estimation methods for multivariate dynamic models. | oktober '76 |
| 30. P. Meulendijks | Keynesiaanse theorieën van handelsliberalisatie. | oktober '76 |
| 31. W. Derks | Structuuranalyse van econometrische modellen met behulp van Grafentheorie. Deel I: inleiding in de Grafentheorie. | september '76 |
| 32. W. Derks | Structuuranalyse van econometrische modellen met behulp van Grafentheorie. Deel II: Formule van Mason. | oktober '76 |
| 33. A. van Schaik | Een direct verband tussen economische veroudering en bezettingsgraadverliezen. | september '76 |
| 34. W. Derks | Structuuranalyse van Econometrische Modellen met behulp van Grafentheorie. Deel III. De graaf van dynamische modellen met één vertraging. | oktober '76 |
| 35. W. Derks | Structuuranalyse van Econometrische Modellen met behulp van Grafentheorie. Deel IV. Formule van Mason en dynamische modellen met één vertraging. | oktober '76 |
| 36. J. Roemen | De ontwikkeling van de omvangsverdeling in de levensmiddelenindustrie in de D.D.R. | oktober '76 |
| 37. W. Derks | Structuuranalyse van Econometrische modellen met behulp van grafentheorie. Deel V. De graaf van dynamische modellen met meerdere vertragingen. | oktober '76 |
| 38. A. van Schaik | Een direct verband tussen economische veroudering en bezettingsgraadverliezen. Deel II: gevoeligheidsanalyse. | december '76 |
| 39. W. Derks | Structuuranalyse van Econometrische modellen met behulp van Grafentheorie. Deel VI. Model I van Klein, statistisch. | december '76 |
| 40. J. Kleijnen | Information Economics: Inleiding en kritiek | november '76 |
| 41. M. v.d. Tillaart. | De spectrale representatie van multivariate zwak-stationaire stochastische processen met discrete tijdparameter. | november '76 |
| 42. W. Groenendaal Th. Dunnewijk | Een econometrisch model van Engeland | december '76 |
| 43. R. Heuts | Capital market models for portfolio selection | september '76 |

| | | |
|--|--|--------------|
| 44. J. Kleijnen en P. Rens | A critical analysis of IBM's inventory package impact. | december '76 |
| 45. J. Kleijnen en P. Rens | Computerized inventory management: A critical analysis of IBM's impact system. | december '76 |
| 46. A. Willemstein | Evaluatie en foutenanalyse van econometrische modellen. Deel I. Een identificatie methode voor een lineair discreet systeem met storingen op input, output en structuur. | januari '77 |
| 47. W. Derks | Structuuranalyse van econometrische modellen met behulp van grafentheorie. Deel VII. Model I van Klein, dynamisch. | februari '77 |
| 48. L. Westermann | On systems of linear inequalities over \mathbb{R}^n . | februari '77 |
| 49. W. Derks | Structuuranalyse van econometrische modellen met behulp van Grafentheorie. Deel VIII. Klein-Goldberger model. | februari '77 |
| 50. W.v. Groenendaal en Th. Dunnewijk | Een econometrisch model van het Verenigd Koninkrijk | februari '77 |
| 51. J. Kleijnen en P. Rens | A critical analysis of IBM's inventory package "IMPACT" | februari '77 |
| 52. J.J.A. Moors | Estimation in truncated parameter-spaces | maart '77 |
| 53. R.M.J. Heuts | Dynamic transfer function-noise modelling (Some theoretical considerations) | december '76 |
| 54. B.B. v.d. Genugten | Limit theorems for LS-estimators in linear regression models with independent errors. | mei '77 |
| 55. P.A. Verheyen | Economische interpretatie in modellen betreffende levensduur van kapitaalgoederen | juni '77 |
| 56. W.v.den Bogaard en J. Kleijnen | Minimizing wasting times using priority classes | juni '77 |
| 57. W. Derks | Structuuranalyse van Econometrische Modellen met behulp van Grafentheorie. Deel IX. Model van landen van de E.E.G. | juni '77 |
| 58. R. Heuts | Capital market models for portfolio selection (a revised version) | juni '77 |
| 59. A.P. Willemstein | Evaluatie en foutenanalyse van econometrische modellen. Deel II. Het Model I van L.R. Klein. | aug. '77 |
| 60. Th. Dunnewijk W. van Groenendaal | An econometric Model of the Federal Republic of Germany 1953-1973 | aug. '77 |
| 61. A. Plaisier A. Hempenius | Slagen of zakken. Een intern rapport over de studieresultaten propedeuse-economie 1974/1975 | aug. '77 |

- | | | |
|-------------------------------|---|-----------|
| 62. A. Hempenius | Over een maat voor de juistheid van voorspellingen | aug. '77 |
| 63. R.M.J. Heuts | Some reformulations and extensions in the univariate Box-Jenkins time series analysis approach (a revised version) | sept. '77 |
| 64. R.M.J. Heuts | Applications of univariate time series modelling of U.S. monetary and business indicator data | sept. '77 |
| 65. A. Hempenius en J. Frijns | Soorten van prijsheteroskedasti- citeit in marktvraagfuncties. | okt. '77 |

Bibliotheek K. U. Brabant



17 000 01059504 0